

Local Averaging Methods

Regression: $X_1, \dots, X_n \in [0, 1]^d$
 $Y_1, \dots, Y_n \in \mathbb{R}$

$$\leadsto \hat{f}(x) = \sum_{i=1}^n w_i(x) Y_i \quad \left. \vphantom{\sum_{i=1}^n} \right\} \text{Local average of the outputs}$$

$$\text{where } \sum_{i=1}^n w_i(x) = 1$$

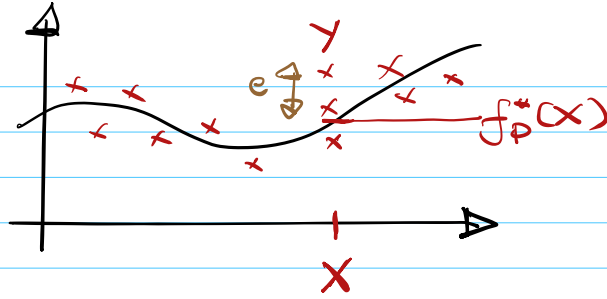
\hookrightarrow if $\left\{ \begin{array}{l} x \text{ close to } X_i : w_i(x) \text{ large} \\ \text{not close} : w_i(x) \text{ small.} \end{array} \right.$

① The Regression problem

$$(X, Y) \sim P$$

$$R_p(f) = \mathbb{E}[(f(X) - Y)^2]$$

Bayes estimator: $f_p^*(x) = \mathbb{E}_p[Y | X=x]$.



Write $Y = f_p^*(x) + e$

error term = $Y - f_p^*(x)$

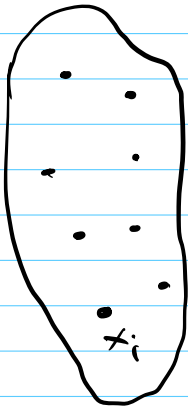
By definition of f_p^* :

$$E_p[e|x] = E[Y|x] - f_p^*(x) = 0.$$

Example:

x_1, \dots, x_n street in a city

y_1, \dots, y_n CO2 concentration



$\hookrightarrow f_p^*(x) =$ Average CO2 concentration at street x .

$$Y = f_p^*(x) + e$$

Noise magnitude may depend on the street x .

\rightarrow Bayes Risk: $R_p^* = R_p(f_p^*)$

$$= E[(f_p^*(x) - Y)^2]$$

$$= E[e^2].$$

What is the excess of risk of a function f ?

$$\mathcal{R}_p(f) = \mathbb{E}_p[(f(X) - Y)^2]$$

$$\begin{aligned}\mathbb{E}[(f(X) - Y)^2 | X] &= \mathbb{E}[(f(X) - f_p^*(X) - e)^2 | X] \\ &= \mathbb{E}[(f(X) - f_p^*(X))^2 | X] + 2 \mathbb{E}[(f(X) - f_p^*(X))e | X]\end{aligned}$$

$$\begin{aligned}&= (f(X) - f_p^*(X))^2 + \mathbb{E}[e^2 | X] \\ &\quad - (f(X) - f_p^*(X)) \underbrace{\mathbb{E}[e | X]}_{=0} = 0\end{aligned}$$

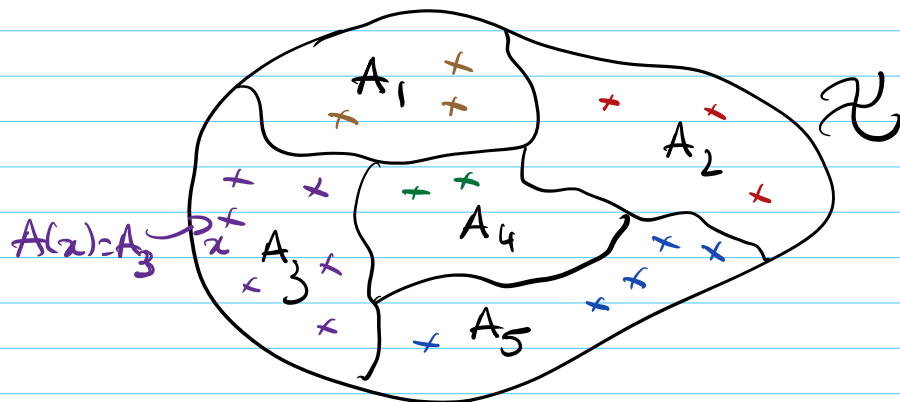
$$= (f(X) - f_p^*(X))^2 + \mathbb{E}[e^2 | X].$$

$$\begin{aligned}\Rightarrow \mathcal{R}_p(f) &= \mathbb{E}[\mathbb{E}[(f(X) - f_p^*(X))^2 | X]] \\ &= \mathbb{E}[(f(X) - f_p^*(X))^2] + \underbrace{\mathbb{E}[\mathbb{E}[e^2 | X]]}_{= \mathbb{E}[e^2]} \\ &= \mathcal{R}_p^*\end{aligned}$$

\Rightarrow Excess of Risk:

$$\left[\begin{aligned}\mathcal{R}_p(f) - \mathcal{R}_p^* &= \mathbb{E}[(f(X) - f_p^*(X))^2] \\ &= \int (f(x) - f_p^*(x))^2 dP_x(x)\end{aligned} \right]$$

② Partition estimator:



Partition $\mathcal{A} = (A_1, \dots, A_5)$ of a set \mathcal{X} .

$$\begin{cases} A_j \cap A_{j'} = \emptyset & j \neq j' \\ \bigcup_{j=1}^5 A_j = \mathcal{X} \end{cases} \quad \mathcal{X} = [0, 1]^d$$

Def: $(X_1, Y_1) \dots (X_n, Y_n)$
 $x \in [0, 1]^d$

$$A(x) = A_j \quad \text{if } x \in A_j.$$

Define

$$w_j(x) = \frac{\mathbb{1}\{X_{j_0} \in A(x)\}}{\sum_{i=1}^n \mathbb{1}\{X_{i_0} \in A(x)\}}$$

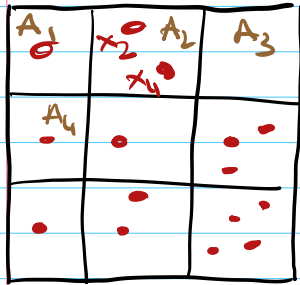
\Rightarrow

$$\hat{f}_{\mathcal{A}}(x) = \sum_{i=1}^n w_i(x) Y_i$$

Partition
estimator

Regression

now let us compute $\hat{f}_A(x)$:



$I_j =$ indexes i such that $x_i \in A_j$
 $n_j =$ size of I_j .

$$n_1 = 1 \quad n_2 = 2 \quad n_3 = 0$$

$$I_2 = \{2, 4\}$$

→ let $x \in A_j$:

$$\begin{aligned} \hat{f}_A(x) &= \sum_{i=1}^n w_i(x) y_i \\ &= \frac{\sum_{i=1}^n \mathbb{1}\{x_i \in A_j\} y_i}{\sum_{i=1}^n \mathbb{1}\{x_i \in A_j\}} \quad n_j \end{aligned}$$

$$\left[\hat{f}_A(x) = \frac{1}{n_j} \sum_{i \in I_j} y_i \right]$$

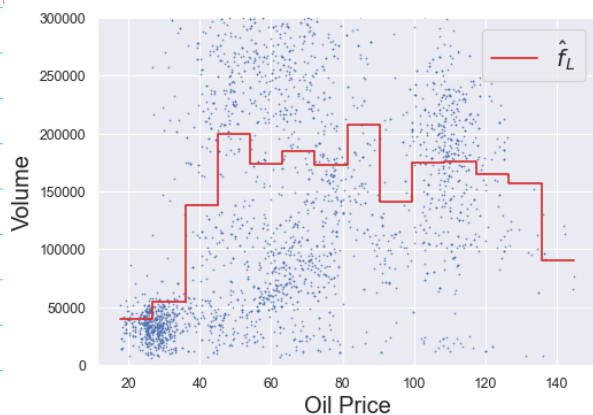
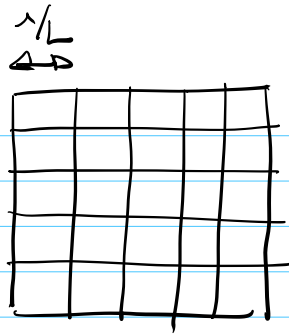


↳ Average of the outputs y_i such that $x_i \in A_j$.

Example:

Cube partition

L^d cubes
of side length $\sim 1/L$.



$d=1$

Is there a
relation between
price and volume
of oil available
on the market?

⇒ What is the excess of risk
of the cube partition estimator?

We assume:

1 f_p^* is α -Lipschitz:

$$\|f_p^*(x) - f_p^*(x')\| \leq \alpha \|x - x'\|$$

2 f_p^* is bounded $|f_p^*(x)| \leq \beta$.

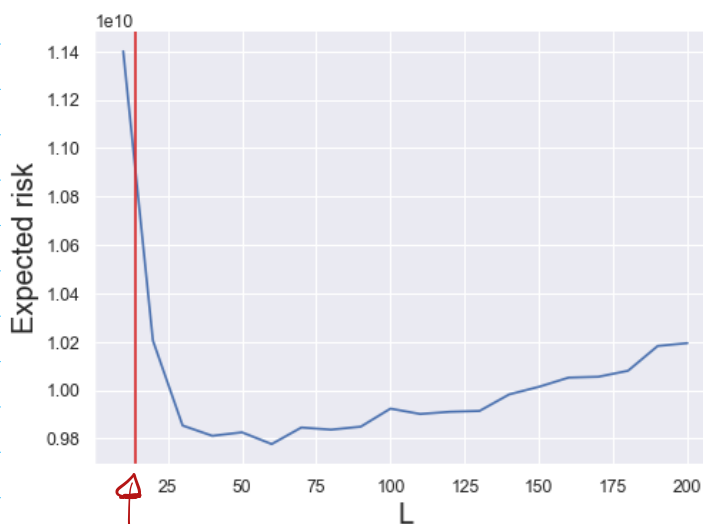
3 Bounded Noise: $|e| \leq \sigma$.

THM: We have

$$\mathbb{E} [\mathcal{R}_p(\hat{f}_L) - \mathcal{R}_p^*] \lesssim \frac{\alpha^2}{L^2} + \frac{\sigma^2 L^d}{n}$$

$$L \approx n^{1/d+2} \Rightarrow \mathbb{E} [\mathcal{R}_p(\hat{f}_L) - \mathcal{R}_p^*] \lesssim n^{-\frac{2}{d+2}}$$

curse of dimensionality



$n^{1/d+2}$

\Rightarrow The theorem gives the order of magnitude of the optimal L .
no cross-validation.

Proof Sketch


$$x \in A_j$$

$$\hat{f}_L(x) - f_P^*(x) = \frac{1}{n_j} \sum_{i \in I_j} y_i - f_P^*(x)$$

$\downarrow = f_P^*(x_i) + e_i$

$$= \frac{1}{n_j} \sum_{i \in I_j} \underbrace{(f_P^*(x_i) - f_P^*(x))}_{\leq \alpha \|x_i - x\|} + \frac{1}{n_j} \sum_{i \in I_j} e_i$$

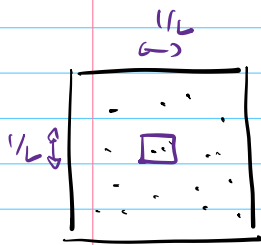
$\leq \alpha \frac{\sqrt{d}}{L}$



$$|\hat{f}_L(x) - f_P^*(x)| \leq \alpha \frac{\sqrt{d}}{L} + \underbrace{\left| \frac{1}{n_j} \sum_{i \in I_j} e_i \right|}$$

Conditionally on I_j , this is an average of n_j r.v.

$$\rightarrow \mathbb{E} \left[\left(\frac{1}{n_j} \sum_{i \in I_j} e_i \right)^2 \mid I_j \right] = \frac{1}{n_j^2} \sum_{i \in I_j} \mathbb{E}[e_i^2 \mid I_j]$$

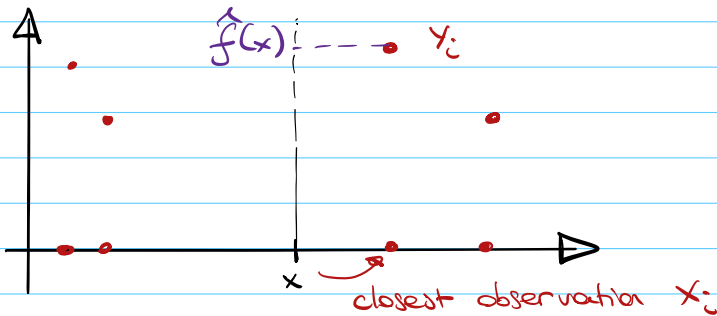


$$n_j \approx n \times L^d$$

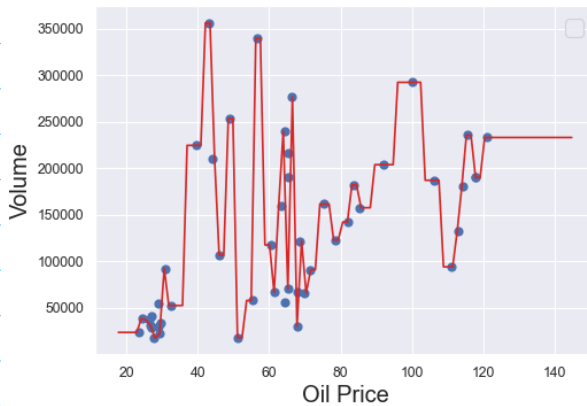
$$\leq \frac{d}{n_j^2}$$

$$\Rightarrow \mathbb{E} \left[|\hat{f}_L(x) - f_P^*(x)|^2 \right] \lesssim \frac{1}{L^2} + \frac{\sigma^2}{n L^d} \quad \square$$

③ Nearest-Neighbor Methods :

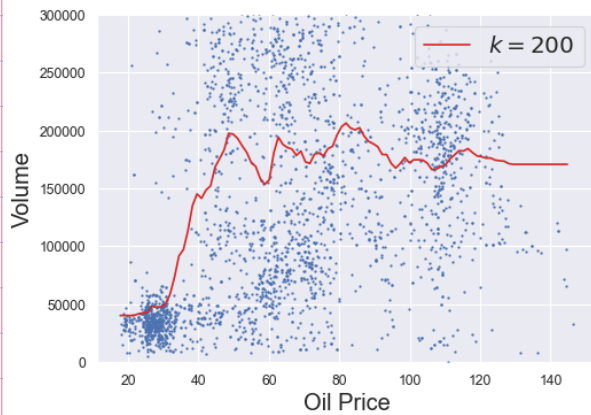


Prediction at x = output y_i of the nearest neighbor x_i .



1 Nearest Neighbor Estimator

Reasonable idea ... but overfitting.



k -Nearest-Neighbor

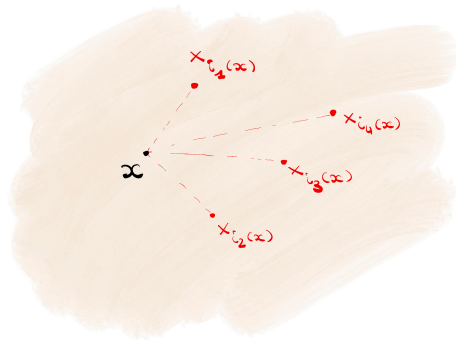


Prediction at x
 = average of the
 k outputs y_i
 corresponding to the
 k NN of x .

Def: $(x_1, y_1) \dots (x_n, y_n) \quad k \geq 1.$
 $x \in [0, 1]^d$

we order the inputs:

$$\|x - X_{i_1(x)}\| \leq \|x - X_{i_2(x)}\| \leq \dots \leq \|x - X_{i_k(x)}\|$$



$\leadsto I_k(x) = \{i_1(x), \dots, i_k(x)\}$ indexes of the k -Nearest-Neighbors
 $w_i(x) = \begin{cases} 1/k & \text{if } i \in I_k(x) \\ 0 & \text{otherwise.} \end{cases}$

$$\begin{aligned} \hat{f}_k^{NN}(x) &= \sum_{i=1}^n w_i(x) y_i \\ &= \frac{1}{k} \sum_{i \in \mathcal{I}_k(x)} y_i \end{aligned}$$

THM: Assume:

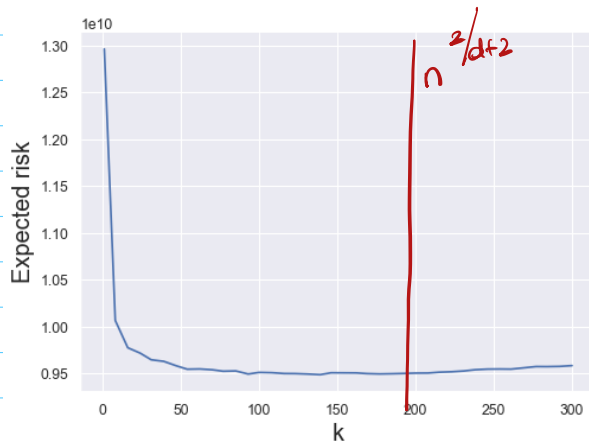
- ① f_p^* is α -Lipschitz
- ② Noise e is bounded: $|e| \leq \sigma$
- ③ X is uniform on $[0,1]^d$.

Then
$$\mathbb{E}[\mathcal{R}_p(\hat{f}_k^{NN}) - \mathcal{R}_p^*] \leq \alpha^2 \left(\frac{k}{n}\right)^{2/d} + \frac{\sigma^2}{k}.$$

For $k \approx n^{2/d+2}$: we obtain

$$\mathbb{E}[\mathcal{R}_p(\hat{f}_k^{NN}) - \mathcal{R}_p^*] \leq n^{-2/d+2}$$

Curse of dimensionality



Select k with cross-validation in practice.

Why does this work?

→ The k NN of x is close to x .
 f_p^* is Lipschitz

$$\Rightarrow \|f_p^*(x) - f_p^*(X_{i_k(x)})\| \leq \alpha \|x - X_{i_k(x)}\|$$

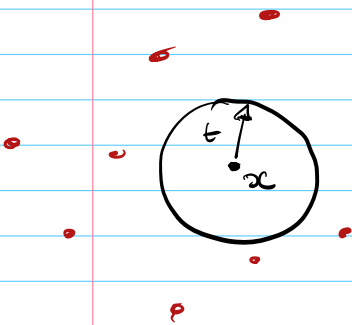
What is the distance between x and its k NN?

$k=1$

$$P(\|x - X_{i_1(x)}\| \geq t) = P(\forall i, \|x - X_i\| \geq t)$$

$$= (1 - P(B(x, t)))^n$$

$$\leq \exp(-n \underbrace{P(B(x, t))}_{\approx t^d})$$



$$\left[E[\|x - X_{i_1(x)}\|^2] \lesssim n^{-2/d} \right]$$

General fact: $E[Z^2] = E\left[\int_0^{+\infty} \mathbb{1}\{t \leq Z^2\} dt\right]$

$$= \int_0^{+\infty} P(Z^2 \geq t) dt$$

$$= 2 \int_0^{+\infty} u P(Z \geq u) du$$

Proof Sketch:

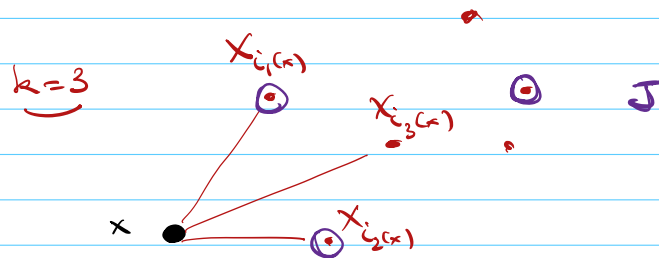
$$|f_k^{NN}(x) - f_p^+(x)| = \left| \frac{1}{k} \sum_{i \in I_k(x)} y_i - f_p^+(x) \right|$$

$$\leq \left(\frac{1}{k} \sum_{i \in I_k(x)} |f_p^+(X_i) - f_p^+(x)| \right)^2 + \left| \frac{1}{k} \sum_{i \in I_k(x)} e_i \right|^2$$

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$$\leq \frac{\alpha^2}{k} \sum_{i \in I_k(x)} \|x - X_i\|^2$$

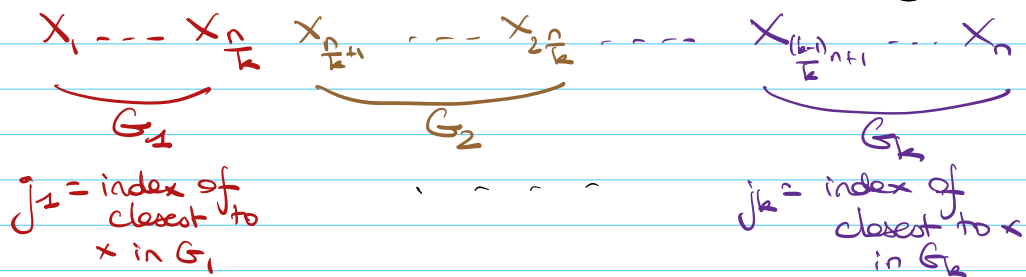
average of
k error terms
 $\Rightarrow \approx \frac{\sigma^2}{k}$



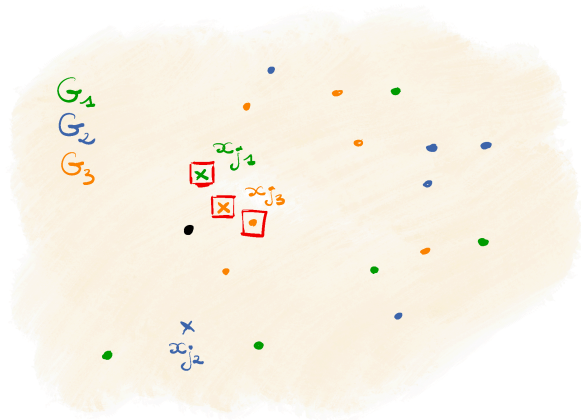
If J is any set of k index:

$$\frac{1}{k} \sum_{i \in I_k} \|x - X_i\|^2 \leq \frac{1}{k} \sum_{i \in J} \|x - X_i\|^2$$

\rightarrow We split the observations in k groups:



$$\rightarrow J = \{j_1, \dots, j_k\}$$



$$\mathbb{E} \left[\frac{1}{k} \sum_{i \in \mathbb{I}_k} \|x - X_i\|^2 \right] \leq \underbrace{\mathbb{E} \left[\frac{1}{k} \sum_{m=1}^k \|x - X_m\|^2 \right]}_{\text{closest to } x \text{ in the group } G_m}$$

$$= \frac{1}{k} \sum_{m=1}^k \mathbb{E} \left[\|x - X_m\|^2 \right]$$

closest to x in the group G_m .

$$\Rightarrow \mathbb{E}[\|x - X_m\|^2] \leq \frac{1}{(n/k)^{2/d}}$$

$$\lesssim \left(\frac{1}{n/k} \right)^{2/d}$$

Conclusion: $\int (f_{\hat{P}}(x) - f_P^+(x))^2 dP(x)$

$$\mathbb{E} \left[\mathcal{R}_P(\hat{f}_k^{NN}) - \mathcal{R}_P^* \right] \lesssim \alpha^2 \left(\frac{1}{n/k} \right)^{2/d} + \frac{1}{k} \rho^2$$

Conclusion :

- * Two local averaging methods:
Partition estimators + k NN
- * Easy to compute and interpret.
- * Depends on one parameter that needs to be tuned.
- * Suffers from the curse of dimensionality.
↳ Rate of CV $n^{-2/d+2}$

Other local averaging method:

Nadaraya - Watson : see Lecture Notes.