

Empirical Risk Minimization

→ What is SUPERVISED LEARNING ?

List of INPUTS:
 $x_1, \dots, x_n \in X$

List of OUTPUTS:
 $y_1, \dots, y_n \in Y$

GOAL: Given a new input x , predict the corresponding y .

Two large families:

CLASSIFICATION

$$Y = \{a, b, c, \dots\}$$

REGRESSION

$$Y = \mathbb{R}$$

Examples:

INPUTS

OUTPUTS

Pictures



Objects

Movie reviews



Review rating

Patient



Is the patient cured?

① Risks and losses:

⚠ There is no single good notion to quantify the quality of a prediction.

Def: A **loss function** is a function

$$l: \mathcal{Y} \times \mathcal{Y} \rightarrow [0, \infty).$$

Examples: ① Binary classification $\mathcal{Y} = \{-1, +1\}$

• Choice 1: $l(y, y') = \begin{cases} 1 & \text{if } y \neq y' \\ 0 & \text{otherwise} \end{cases}$

• Choice 2: $l(y, y') = \begin{cases} a & \text{if } y = 1 \text{ and } y' = -1 \\ b & \text{if } y = -1 \text{ and } y' = 1 \\ 0 & \text{if } y = y' \end{cases}$

→ Choice 2 makes sense (for instance) in medical settings, where it is a more serious mistake to predict that a patient is not sick ($y' = -1$) whereas they are ($y = +1$) than the opposite.

② Regression $Y = \mathbb{R}^D$ $y = (y_1, \dots, y_D)$ $y' = (y'_1, \dots, y'_D)$

• L_∞ norm $\|y' - y\|_\infty = \max_{i=1 \dots D} |y_i - y'_i|$

• L_p norm $\|y' - y\|_p = \left(\sum_{i=1}^D |y_i - y'_i|^p \right)^{1/p}$

• Weighted L_p norm $\left(\sum_{i=1}^D w_i |y_i - y'_i|^p \right)^{1/p}$

Goal: Find a predictor $f: X \rightarrow Y$ such that, $\ell(f(x), y)$ is small on New samples $(x'_1, y'_1), \dots, (x'_m, y'_m)$.

TRAINING
SAMPLES
 (x_1, y_1)
...
 (x_n, y_n)

→ Find a predictor f .

NEW SAMPLES
 (x'_1, y'_1) ... (x'_m, y'_m)

We want $\ell(f(x'_i), y'_i)$ to be small on average.
(i.e. good predictions on new samples)

Example: Electric consumption Forecasting

Training samples: $y_i =$ Electric consumption
on Day i

$x_i =$ characteristics of Day i
(weather, day of the week,
etc.)

TRAIN
A MODEL



Testing sample: Predict tomorrow's consumption
based on tomorrow's weather, etc.

ASSUMPTION: The observations $(x_1, y_1) \dots (x_n, y_n)$
are i. i. d. with law P .

independent identically distributed.

Def Given a function $f: X \rightarrow Y$, the
P-risk of f is defined as:

$$R_P(f) := \mathbb{E}_P[\ell(f(X), Y)]$$

Theorem: The P-risk is minimized for the

Bayes predictor f_P^* defined by

$$f_P^*(x) \in \operatorname{argmin}_{z \in Y} \mathbb{E}_P[\ell(y, z) | X=x].$$

proof: Let $\Psi(x, z) = \mathbb{E}_p[\ell(Y, z) | X=x]$

$\leadsto \Psi(x, z) \geq \Psi(x, f_p^*(x))$ (by definition)

$$R_p(f) = \mathbb{E}_p[\ell(Y, f(X))] \stackrel{(*)}{=} \mathbb{E}_p[\Psi(X, f(X))]$$

$$\geq \mathbb{E}_p[\Psi(X, f_p^*(X))]$$

$$\stackrel{(*)}{=} \mathbb{E}_p[\ell(Y, f_p^*(X))] = R_p(f_p^*)$$

(*) LAW OF TOTAL EXPECTATION

$$\mathbb{E}[A] = \mathbb{E}[\mathbb{E}[A|X]]$$



Examples:

① Binary classification:

$$\ell(y, y') = \begin{cases} 1 & \text{if } y \neq y' \\ 0 & \text{otherwise} \end{cases}$$

$$\mathbb{E}_p[\ell(Y, z) | X=x] = P(Y \neq z | X=x)$$

$$z = -1 \text{ or } +1 \quad = 1 - P(Y=z | X=x)$$

\rightarrow if $\eta(x) = P(Y=1 | X=x) \geq \frac{1}{2}$,

then $f_p^*(x) = 1$.

Otherwise, $f_p^*(x) = -1$.

② Quadratic loss $\mathcal{Y} = \mathbb{R}$ $l(y, y') = (y - y')^2$

Question: A random variable.
What is the minimum of

$$a \mapsto \mathbb{E}[(A - a)^2] = \mathbb{E}[A^2] - 2\mathbb{E}[A]a + a^2$$

$$\partial_a = 2(a - \mathbb{E}[A]) \quad \text{so } a = \mathbb{E}[A]$$

$$\rightarrow f_p^*(x) = \mathbb{E}_p[Y | X = x]$$

② Empirical risk

\mathbb{P} is unknown $\Rightarrow f_p^*$ is unknown

GOAL: Approximate f_p^* using the observations

$$(x_1, y_1), \dots, (x_n, y_n)$$

Def: The **empirical risk** of a function f is

$$\left[\mathcal{R}_n(f) = \frac{1}{n} \sum_{i=1}^n l(f(x_i), y_i) \right]$$

LAW OF LARGE NUMBERS

$$R_n(f) = \frac{1}{n} \sum_{i=1}^n \ell(f(x_i), y_i)$$

$$\downarrow n \rightarrow +\infty$$

$$R_p(f) = \mathbb{E}_p[\ell(f(x), y)]$$

\Rightarrow Minimizing $R_n(f) \approx$ Minimizing $R_p(f)$

Def: Let \mathcal{F} be a set of functions from $x \rightarrow y$.

The empirical risk minimizer of \mathcal{F} is

$$\left[\hat{f}_{\mathcal{F}} \in \operatorname{argmin}_{f \in \mathcal{F}} R_n(f) \right]$$

Examples:

① Linear Regression: $x = \mathbb{R}^d$ $y = \mathbb{R}$

$$\ell(y, y') = (y - y')^2$$

$$\mathcal{F}_{\text{lin}} = \left\{ f_{\theta} : x \mapsto \langle x, \theta \rangle : \theta \in \mathbb{R}^d \right\}$$

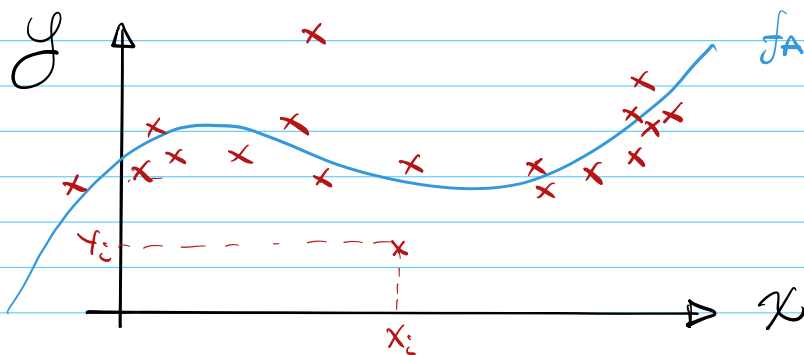
$$\begin{aligned}
 R_n(f_\theta) &= \frac{1}{n} \sum_{i=1}^n (y_i - f_\theta(x_i))^2 \\
 &= \frac{1}{n} \sum_{i=1}^n (y_i - \langle x_i, \theta \rangle)^2 \\
 &= \frac{1}{n} \|y - X\theta\|^2
 \end{aligned}
 \left. \vphantom{\begin{aligned} R_n(f_\theta) &= \frac{1}{n} \sum_{i=1}^n (y_i - f_\theta(x_i))^2 \\ &= \frac{1}{n} \sum_{i=1}^n (y_i - \langle x_i, \theta \rangle)^2 \\ &= \frac{1}{n} \|y - X\theta\|^2 \end{aligned}} \right\} \text{LINEAR REGRESSION}$$

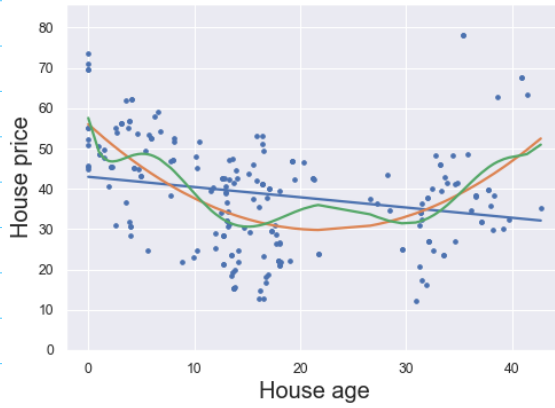
$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_i \\ \vdots \\ y_n \end{pmatrix} \quad X = \begin{bmatrix} x_1 & \vdots & x_n \end{bmatrix}$$

② Polynomial Regression

$$X = \mathbb{R} \quad Y = \mathbb{R} \quad \ell(y, y') = |y - y'|^2$$

$$\mathcal{F}_d = \left\{ x \mapsto \underbrace{\sum_{i=0}^d a_i x^i}_{= f_A} : \underbrace{a_0 \dots a_d}_{= A \in \mathbb{R}^{d+1}} \in \mathbb{R} \right\}$$

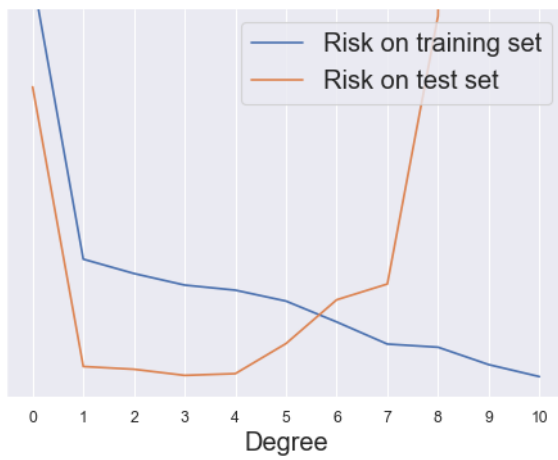




— $d=1$

— $d=2$

— $d=8$



→ Let us explain this phenomenon.



Training set $(x_1, y_1) \dots (x_n, y_n)$

⇒ Find prediction \hat{f}_F **DEPENDENT** on the Training set.

⇒ We can NOT apply the Law of Large numbers to say that $R_p(\hat{f}_F) \approx R_n(\hat{f}_F)$

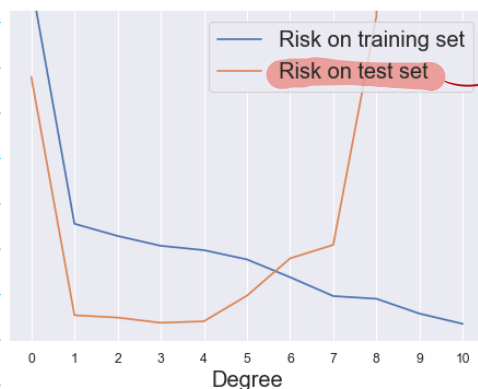
$$= \frac{1}{n} \sum_{i=1}^n \underbrace{\ell(\hat{f}_F(x_i), y_i)}_{\text{not indep!!}}$$

⇒ On the testing sample $(x'_1, y'_1) \dots (x'_m, y'_m)$ **INDEPENDENT** from the training sample,

LLN CONDITIONALLY ON THE TRAINING SAMPLE



$$R_p(\hat{f}_F) \approx \frac{1}{m} \sum_{i=1}^m \underbrace{\ell(\hat{f}_F(x'_i), y'_i)}_{\text{conditionally independent.}}$$



Good Approximation of $R_p(\hat{f}_F)$

Bound on the Estimation Error:

$$\begin{aligned} \mathcal{R}_P(\hat{f}_S) - \inf_{f \in \mathcal{F}} \mathcal{R}_P(f) &= \mathcal{R}_P(\hat{f}_S) - \mathcal{R}_P(f_S^*) \\ &= (\mathcal{R}_P(\hat{f}_S) - \mathcal{R}_n(\hat{f}_S)) \\ &\quad \left[+ (\mathcal{R}_n(\hat{f}_S) - \mathcal{R}_n(f_S^*)) \right] \leq 0 \\ &\quad + (\mathcal{R}_n(f_S^*) - \mathcal{R}_P(f_S^*)) \\ &\leq \underbrace{\sup_{f \in \mathcal{F}} (\mathcal{R}_P(f) - \mathcal{R}_n(f))}_{\text{uniform the deviation between empirical risk and P-risk}} + (\mathcal{R}_n(f_S^*) - \mathcal{R}_P(f_S^*)) \end{aligned}$$

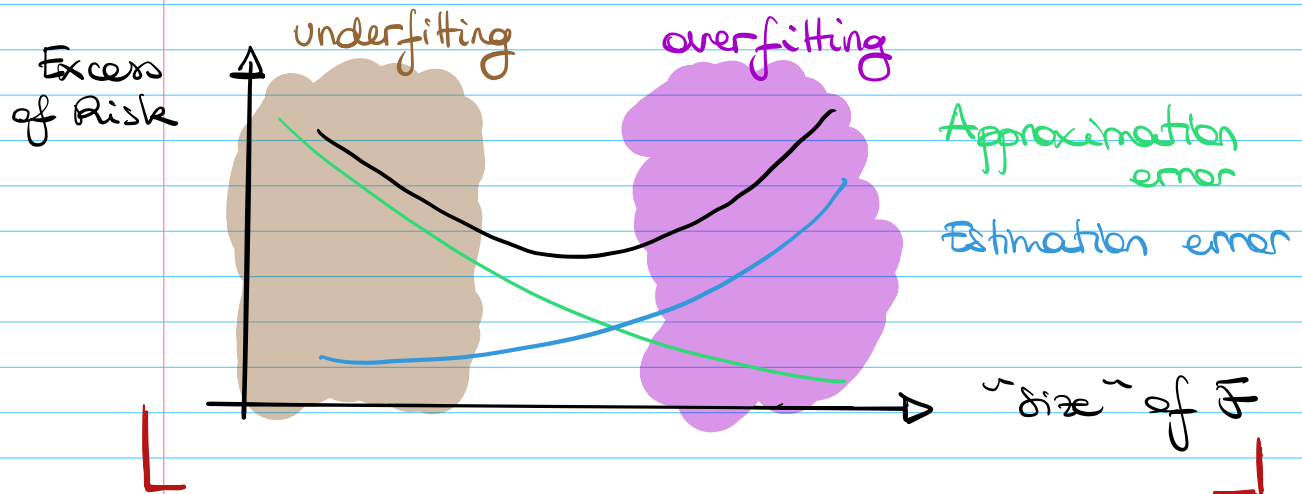
uniform the deviation between empirical risk and P-risk.

→ CENTRAL LIMIT THEOREM:

$$\begin{aligned} \mathcal{R}_P(f) - \mathcal{R}_n(f) &= \frac{1}{n} \sum_{i=1}^n (E[\ell(f(x_i), Y_i)] - \ell(f(x_i), Y_i)) \\ &\approx \frac{1}{\sqrt{n}} \times \text{Gaussian} \end{aligned}$$

↳ But what about $\sup_{f \in \mathcal{F}} (\mathcal{R}_P(f) - \mathcal{R}_n(f))$?

The Picture



③ Bound on the estimation error in Binary Classification

$$\mathcal{Y} = \{-1, 1\} \quad \ell(y, y') = \mathbb{1}\{y \neq y'\}$$

GOAL: Bound $\mathbb{E} \left[\sup_{f \in \mathcal{F}} (R_p(f) - R_n(f)) \right]$

CASE 1: Finite Number of Predictors

$$\mathcal{F} = \{f_1, \dots, f_k\}$$

THEOREM (Maximal Inequality)

Let Z_1, \dots, Z_k be random variables with

$$\forall d > 0, \quad \mathbb{E}[e^{dZ_j}] \leq e^{d^2\sigma^2/2} \quad (\text{subgaussianity condition})$$

Then,

$$\left[\mathbb{E}\left[\max_{j=1, \dots, k} Z_j\right] \leq \sigma\sqrt{2\log k} \right]$$

proof:

$$\max_{j=1, \dots, k} e^{dZ_j} \leq \sum_{j=1}^k e^{dZ_j} \quad (*)$$

$$\Rightarrow \mathbb{E}\left[\max_{j=1, \dots, k} Z_j\right] = \frac{1}{d} \mathbb{E}\left[\log\left(\max_{j=1, \dots, k} e^{dZ_j}\right)\right]$$

$$\uparrow \frac{1}{d} \log(e^{dx}) = x$$

JENSEN

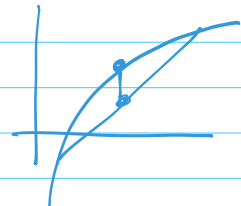
$$\leq \frac{1}{d} \log\left(\mathbb{E}\left[\max_{j=1, \dots, k} e^{dZ_j}\right]\right)$$

$$\stackrel{(*)}{\leq} \frac{1}{d} \log\left(\mathbb{E}\left[\sum_{j=1}^k e^{dZ_j}\right]\right)$$

$$\underbrace{\sum_{j=1}^k \mathbb{E}[e^{dZ_j}]} \leq k e^{d^2\sigma^2/2}$$

$$\leq \frac{\log k}{d} + \frac{1}{d} \frac{d^2\sigma^2}{2}$$

no Choose $d = \frac{\sqrt{2\log k}}{\sigma}$. ▣



Recall

$$\begin{aligned} & \mathbb{E} \left[R_P(\hat{f}_F) - \inf_{f \in F} R_P(f) \right] \\ & \leq \mathbb{E} \left[\sup_{f \in F} (R_P(f) - R_n(f)) \right] + \mathbb{E} \left[(R_n(f_F^*) - R_P(f_F^*)) \right] \\ & = \frac{1}{n} \sum_{i=1}^n \mathbb{1} \{ f_F^*(x_i) \neq y_i \} - \mathbb{E}_P \left[\mathbb{1} \{ f_F^*(x) \neq y \} \right] \\ & \Rightarrow \mathbb{E} [\text{green oval}] = 0 \end{aligned}$$

See notes: $R_P(f) - R_n(f)$ is $\frac{1}{\sqrt{n}}$ -subgaussian.

BOUND ON THE ESTIMATION ERROR

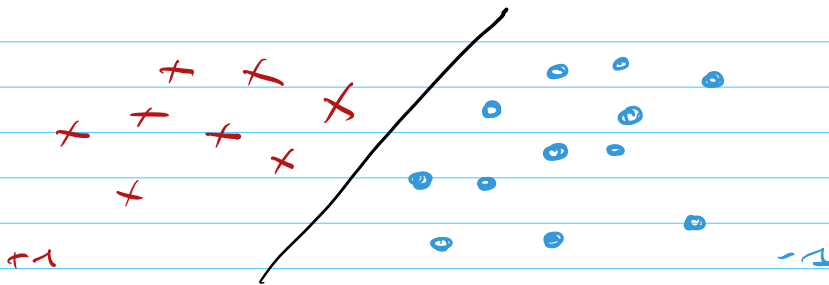
$$\Rightarrow \left[\mathbb{E} \left[R_P(\hat{f}_F) - \inf_{f \in F} R_P(f) \right] \leq \sqrt{\frac{2 \log k}{n}} \right]$$

size of F .

CASE 2: With VC-dimension

What if F is infinite?

ex: $\mathcal{F}_{\text{lin}} = \{\text{hyperplane classifiers}\}$



Even if \mathcal{F} is infinite, there is only a finite number of classifications.

$$\mathcal{C}_{\mathcal{F}}(x_1, \dots, x_n) = \{ (f(x_1), \dots, f(x_n)) : f \in \mathcal{F} \}$$
$$\subseteq \{-1, +1\}^n$$

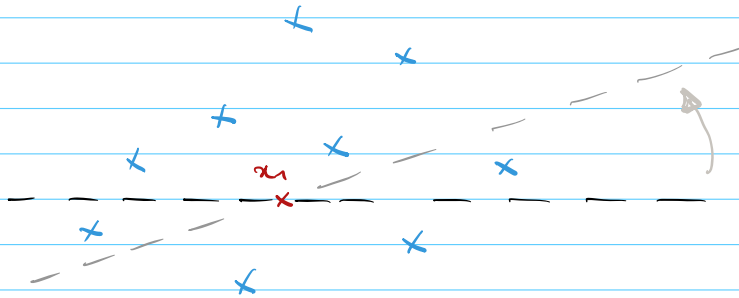
of size $|\mathcal{C}_{\mathcal{F}}(x_1, \dots, x_n)| \leq 2^n$

⇒ Actually the size of \mathcal{F} can be replaced by $|\mathcal{C}_{\mathcal{F}}(x_1, \dots, x_n)|$!

THEOREM :

$$\mathbb{E} \left[\sup_{f \in \mathcal{F}} (R_p(f) - R_n(f)) \right] \leq 2 \mathbb{E} \left[\sqrt{\frac{2 \log |\mathcal{C}_{\mathcal{F}}(x_1, \dots, x_n)|}{n}} \right]$$

Example: $\mathcal{H} = \mathcal{H}_{\text{lin}}$ in \mathbb{R}^2



Rotation around x_1 : $\leq n$ different classifications
Rotation around x_2 : $\leq n$ different classifications
⋮
⋮
Rotation around x_n : $\leq n$ different classifications

⇒ At most n^2 different classifications!

$$\mathcal{N}_{\mathcal{H}}(x_1, \dots, x_n) \leq n^2$$

$$\Rightarrow \mathbb{E}[\text{estimation error}] \leq 4 \sqrt{\frac{\log n}{n}}$$

And for a general \mathcal{F} ?

- We say that \mathcal{F} shatters (x_1, \dots, x_n) if $N_{\mathcal{F}}(x_1, \dots, x_n) = 2^n$.

→ Vapnik-Chervonenkis

- The VC dimension of \mathcal{F} is the largest n such that there exists a configuration (x_1, \dots, x_n) of n points being shattered by \mathcal{F} .

→ $VCC(\mathcal{F})$

For $n \leq VCC(\mathcal{F})$, we cannot bound

$N_{\mathcal{F}}(x_1, \dots, x_n)$ meaningfully



What if $n > VCC(\mathcal{F})$?



MAGICAL RESULT

SAUER LEMMA: if $n > VCC(\mathcal{F})$,

$$\log N_{\mathcal{F}}(x_1, \dots, x_n) \leq VCC(\mathcal{F}) \log \left(\frac{en}{VCC(\mathcal{F})} \right)$$

→ if $n > 2VCC(\mathcal{F})$,

$$N_{\mathcal{F}}(\dots) \lesssim n^{VCC(\mathcal{F})} \ll 2^n !!$$

THEOREM: if $n > 2VCCF$

$$E[R_p(\hat{f}_F) - \inf_{f \in F} R_p(f)] \leq 2 \sqrt{\frac{2VCCF}{n} \log\left(\frac{en}{VCCF}\right)}$$

Estimation error in binary classification $\leq \sqrt{\frac{VCCF}{n}}$

TAKE-HOME MESSAGES:

- The quality of a prediction is measured by a loss l .

$$P\text{-Risk} = E_P[\text{loss}]$$

- The best theoretical predictor is the **Bayes predictor**.
 \rightarrow cannot be computed
- A strategy to find a good predictor is to minimize the **empirical risk** on a model \mathcal{F} .
- The "size" of \mathcal{F} has to be properly chosen to avoid

UNDERFITTING and **OVERFITTING**

Large approximation error

\hookrightarrow The model \mathcal{F} is too simple to capture the complexity of the dataset

Large estimation error

\hookrightarrow in binary classification, bounded by $\approx \sqrt{\frac{V(\mathcal{F})}{n}}$

Large if \mathcal{F} is too complex.